

Stochastic processes, stationarity and autocorrelation

Individual stochastic processes

- A stochastic process is an ordered sequence of random variables: $\dots y_0, y_1, \dots, y_t, y_{t+1}, \dots$ or $\{y_t\}$ for short.
- Each panel member is characterised by a different stochastic process, so add subscript i : $\{y_{it}\}$
- Assume two things:
 - (A1) individuals i and j initially sampled at random at (say) time $t = 0$
 \Rightarrow joint distribution of y_{i0} and y_{j0} is $f_0(y_{i0}) \times f_0(y_{j0})$
 - (A2) after selection, there is no cross-individual influence \Rightarrow joint distribution of $(y_{it}, y_{jt} | y_{i0}, y_{j0})$ is $f_1(y_{it} | y_{i0}) \times f_1(y_{jt} | y_{j0})$
- Implies that the whole sequence $\{y_{it}\}$ for individual i is statistically independent of the sequence $\{y_{jt}\}$ for any other individual j

Stationarity

- A stochastic process $\{y_t\}$ is stationary (specifically 2nd-order or covariance-stationary) if:
 - the mean, $E(y_t)$, is the same for all periods t
 - the variance, $Var(y_t)$, is the same for all periods t
 - each covariance, $Cov(y_t, y_s)$, only depends on the time separation $|t-s|$, not on the particular times t and s at which y is observed
- Stationarity \Rightarrow in any sequence of consecutive periods we happen to observe, the process $\{y_t\}$ will tend to look much the same - at least in terms of its first and second moments
- Most panel data don't follow stationary processes - e.g. earnings, BMI, happiness, etc. tend to follow clear trends in level and volatility through the lifecycle
- But, stationary processes are useful as building blocks to construct realistic models

Five important stochastic processes

- *White noise* $\{\varepsilon_t\}$ is a sequence of independent random variables, each with the same mean 0 and variance σ_ε^2
- An *Equi-correlated* or *exchangeable* process can be expressed as the sum of a time-invariant random variable u and white noise ε_t

$$y_{it} = u_i + \varepsilon_{it}$$
- A k th-order *autoregressive process*, or AR(k), involves feedback from k past values and a white noise innovation:

$$y_{it} = \rho_1 y_{it-1} + \dots + \rho_k y_{it-k} + \varepsilon_{it}$$
- A k th-order *moving average process*, or MA(k), is an average of k consecutive values from a white noise process:

$$y_{it} = \varepsilon_{it} + \theta_1 \varepsilon_{it-1} + \dots + \theta_k \varepsilon_{it-k} + \varepsilon_{it}$$
- A *random walk* is a **non-stationary** process that cumulates white noise innovations:
 - without drift: $y_{it} = y_{it-1} + \varepsilon_{it}$
 - with drift: $y_{it} = \alpha + y_{it-1} + \varepsilon_{it}$

Autocovariances and autocorrelations

For a stationary stochastic process $\{y_{it}\}$, define the j th-order autocovariance, $c(j)$, and autocorrelation, $\rho(j)$ as:

$$c(j) = \text{cov}(y_{it}, y_{it-j})$$

$$r(j) = \text{corr}(y_{it}, y_{it-j})$$

For non-stationary processes, $c(j)$ and $r(j)$ may depend on t .

Non-stationary example: random walk

$$y_{it} = \varepsilon_{it} + \varepsilon_{it-1} + \dots + \varepsilon_{i1} \quad (\text{initial condition: } y_{i0} = 0)$$

Variance: $\sigma_y^2 = t\sigma_\varepsilon^2 \rightarrow \infty$ as t increases

Autocovariance: $c(j) = (t-j)\sigma_\varepsilon^2$

Autocorrelation: $r(j) = (t-j) / [t(t-j)]^{1/2} \rightarrow 1$ as t increases, for all j

Autocovariances and autocorrelations

An equicorrelated process, $y_{it} = u_i + \varepsilon_{it}$, has variance and serial correlation:

$$\sigma_y^2 = \sigma_u^2 + \sigma_\varepsilon^2$$

$$c(j) = \sigma_u^2$$

$$r(j) = \sigma_u^2 / (\sigma_u^2 + \sigma_\varepsilon^2) \Rightarrow \text{doesn't decay with increasing } j$$

An AR(1) process, $y_{it} = \rho y_{it-1} + \varepsilon_{it}$, has variance and serial correlation:

$$\sigma_y^2 = \sigma_\varepsilon^2 / (1 - \rho^2)$$

$$c(j) = \sigma_y^2 \rho^j$$

$$r(j) = \rho^j \Rightarrow \text{geometric decay}$$

An MA(1) process, $y_{it} = \varepsilon_{it} + \theta \varepsilon_{it-1}$, has variance and serial correlation:

$$\sigma_y^2 = \sigma_\varepsilon^2(1 + \theta^2)$$

$$c(1) = \sigma_\varepsilon^2 \theta; c(j) = 0 \text{ for all } j > 1$$

$$r(1) = \theta / (1 + \theta^2); r(j) = 0 \text{ for all } j > 1 \Rightarrow \text{abrupt decay}$$

NB $r(1)$ can never exceed 0.5 for an MA(1) process

Monte Carlo simulations

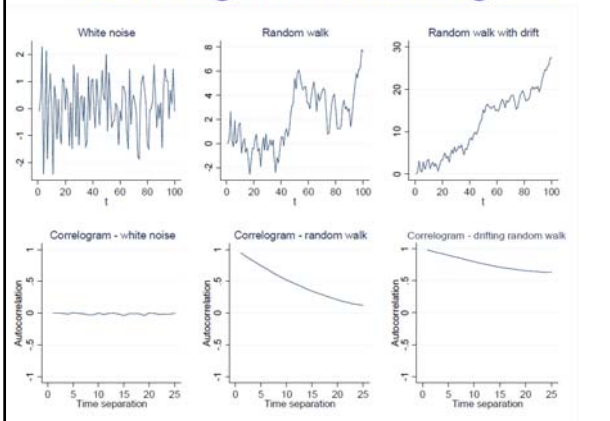
Six processes:

- White noise: $y_{it} = \varepsilon_{it}$
- Equi-correlated: $y_{it} = u_i + \varepsilon_{it}$
- AR(1): $y_{it} = 0.75 y_{it-1} + \varepsilon_{it}$
- MA(1): $y_{it} = \varepsilon_{it} + 0.9 \varepsilon_{it-1}$
- Random walk: $y_{it} = y_{it-1} + \varepsilon_{it}$
- Random walk with drift: $y_{it} = 0.2 + y_{it-1} + \varepsilon_{it}$

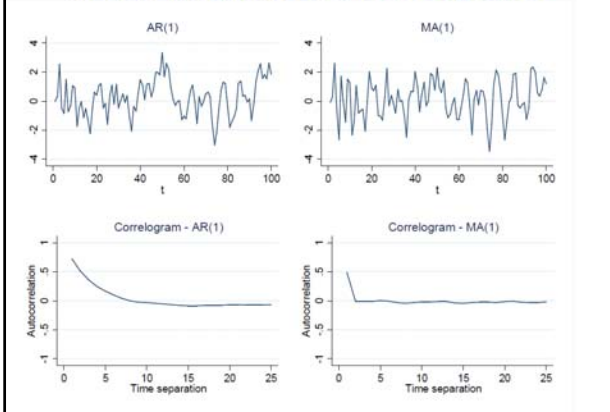
Method:

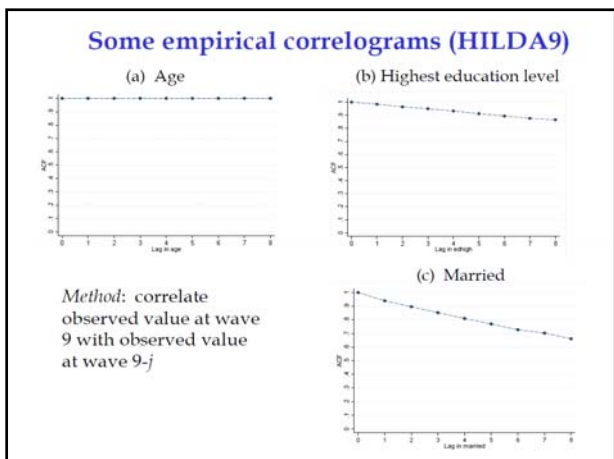
- generate pseudo-random $N(0,1)$ values for $u_i, \varepsilon_{i1} \dots \varepsilon_{i100}$
- from initial condition $y_{i0} = 0$, calculate sequence of $y_{i1} \dots y_{i100}$
- scatter plot of $\{y_{it}\}$ for $i = 1$
- estimate and plot the first 25 elements of the correlogram

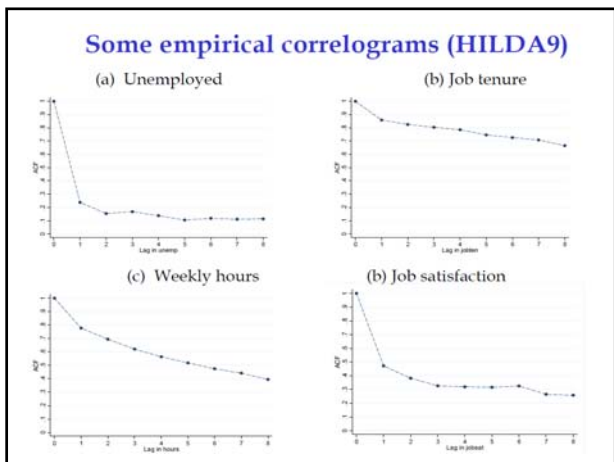
Scatter diagrams and correlograms

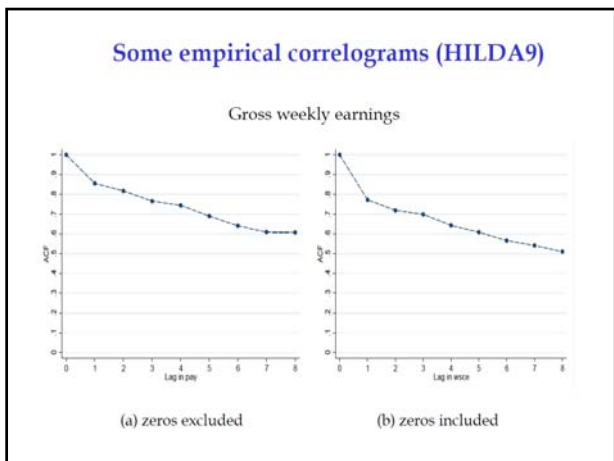


Scatter diagrams and correlograms









Typical properties of real panel data

- Correlograms usually display slow decay, high persistence
- eg: autocorrelations at a separation of 8 years:
 - 1.0 (trivially) age, gender, etc.
 - 0.8-1.0 educational attainment
 - 0.6-0.8 marital status, job tenure
 - 0.4-0.6 weekly hours, earnings
 - 0.2-0.4 job satisfaction
 - < 0.2 unemployed
- Persistence \Rightarrow the total no. of observations, nT , greatly exaggerates the amount of information in a panel
- Should we assume that conditions external to the individual (i.e. the covariates) capture persistence adequately, or does behaviour display inherent persistence (e.g. habit, inertia)?
